

## Appendix A: Clinical data on cases of testicular ascent

*Clinical data on cases of testicular ascent*

References	No. Cases	Mean Age (yrs)	Position (at surgery)				No. Hernias/ Testes	No. Side (unilat/bilat)
			Prescrotal/ High Scrotal	Superficial Inguinal Pouch	Inguinal	Abdominal		
Myers and Officer <sup>10</sup>	7	6.6	—	—	—	—	3/4	
Atwell <sup>28</sup>	10	9.4	7	1	2	—	10/11	
Schiffer et al <sup>29</sup>	3	6.7	—	2	—	—	3/0	
Belman <sup>30</sup>	6	—	1	4	1	—	3/6	
Robertson et al <sup>31</sup>	13	8.1	—	—	—	—	10/13	
Eardley et al <sup>32</sup>	34	7.5	15	19	3	—	24/39	
Mayr et al <sup>33</sup>	19	7.0	5	9	13	—	14/27	
Rabinowitz and Hulbert <sup>34</sup>	21	7.2	2	19	2	—	12/23	
Gracia et al <sup>35</sup>	36	7.0	8	5	32	1	18/46	
Clarnette et al <sup>36</sup>	25	7.6	—	—	—	—	10/33	
Rusnack et al <sup>37</sup>	91	7.4	79	11	—	1	39/91	
Total No. (%)	265		117 (48)	70 (29)	53 (22)	2	140/289 (48)	

\* As published in *Journal of Urology*

Barthold JS and Gonzalez R: The epidemiology of congenital cryptorchidism, testicular ascent and orchiopexy. *J Urol* 2003; 170: 2396.

## Appendix B: Evidence table of studies examining the association of CAG repeats in the AR gene and cryptorchidism

Sasagawa2000	Ferlin2005	Silva-Ramos2006	Radpour2007
<b>Study characteristics</b>			
Tokyo, Japan Case-control	Padova, Italy Case-control, prospectively recruited.	Porto, Portugal Case-control, prospectively recruited from two hospitals.	Tehran, Iran Case-control
<b>Sample characteristics</b>			
48 Japanese Pts w/crypto (age 1-32, mean 13 yrs); 17 BL, 29 UL UDT and 2 w monorchia). Controls: 100 males w/ proven fertility	105 ex-cryptorchid men (55 UL; 50 BL) who presented no other obvious causes of testicular damage. Age 23-42 yrs, age at orchidopexy 1-12 years . Precise location of testes at surgery could not be determined in all cases.  115 fertile non-cryptorchid controls whose wives were pregnant.	42 cryptorchid boys (age 3 -77; mean 20.9 yrs); 7 BL, 35 UL (6 w/ clinically patent processus vaginalis (PPV) in the contra-lateral; 3 UL ectopic; 2 UL intra-abdominal). Six w/ family hx of crypto. (Cases should not present any other genital malformation).  31 controls (hospital staff) w/o personal or family hx of genital abnormalities; age 22-69 yrs, mean 42 yrs.	76 unrelated Iranian males w/ crypto (27 UL; 49 BL) w/o spontaneous descent.  190 healthy fertile controls.  All studied patients were 46,XY males. Patients with visible cytogenetic aberrations were excluded.
<b>Measurement of CAG repeats</b>			
Genomic DNA from peripheral leukocytes amplified by PCR (30 cycles). Repeat length analysis by GeneScan.	Genomic DNA extracted from peripheral blood leuco-cytes. AR exon 1 amplified in two different PCR reactions under the same standard conditions with 8% dimethylsulphoxide and same cycle.	Genomic DNA extracted from blood samples stored at -20°C & amplified by PCR.	Genomic DNA from peripheral blood lymphocytes was amplified by PCR in 2 different reactions. Genotyping done blinded to case-control status.
<b>Cryptorchidism ascertainment</b>			
Not indicated.	Cohort of men previously orchidopexied.	Sole inclusion criterion was the presence of crypto.	Ultrasound examination of the testes.
<b>CAG repeats among cryptorchidism cases (mean ± SD; second line range and median)</b>			
23.4 ± 0.48 (overall) Range: 16 – 32; median: 23	21.8 ± 2.9 (overall 105 crypto) Range: 12 – 29; median: 22	22.4 ± 3.2 (overall) (p-v=0.206) Range: 15 – 29; median: 22.5	21.9 ± 2.5 (overall) Range: 17 – 30; median: 21
23.7 ± 0.64 (31 UL) Range: 18 – 32; median: 24	21.8 ± 3.0 (55 UL) Range: 12 – 29; median: 21	22.1 ± 3.3 (35 UL) (p-v=0.474) Range: 15 – 29; median: 21	21.5 ± 0.8 (27 UL) Range: 17 – 29; median: 21
23.8 ± 0.69 (17 BL) Range: 16 – 28; median: 23;	21.8 ± 2.7 (50 BL) Range: 13 – 27; median: 22	24.7 ± 1.8 (7 BL) (p-v=0.019) Range: 21 – 26; median: 25.5	22.4 ± 1.6 (49 BL) Range: 19 – 30; median: 22
(SE rather than SD provided)		25.2 ± 2.9 (6 UL w/ PPV) Range: 20 – 29; median: 25	
<b>CAG repeats among controls</b>			
23.5 ± 0.29 Range: 15 – 32; median: 23	21.6 ± 3.3 Range: 9 – 31; median: 22	21.5 ± 2.7 Range: 16 – 26; median: 22	21.8 ± 2.8 Range: 16 – 28; median: 21

## Appendix C - Evidence table of studies examining the association of ESR1 SNPs and cryptorchidism

Yoshida2005	Galan2007	Wang2008												
<b>Study characteristics</b>														
Tokyo, Japan, 1995-2003 Case-control	Florence, Italy Case-control	USA (Wilmington, DE) Case-control												
<b>Sample characteristics</b>														
63 Japanese cryptorchid males, aged 1–13 yr, (median, 5.0 yrs) and 47 control males, aged 4–12 yr (median, 7.5 yrs) 47 UL: inguinal region in 39; abdominal cavity in 7; 1 vanishing testis. 16 BL: both testes inguinal region in 15; abdominal cavity in 1  Individuals recruited from the urban or suburban area of Tokyo or Yamagata City. They were free of particular residential environments, specific dietary habits, and intake of drugs with hormonal effects.	118 cryptorchid males of Italian origin (Central Italy); 71 UL, 47 BL  The Italian control group includes 168 men from Central Italy with normal sperm parameters and no hx of cryptorchidism.	152 cryptorchid males who had non-syndromic UL (120) or BL (32) UDT, and had no other discernible congenital abnormalities UDT located in the prescrotal region, external ring, or superficial inguinal pouch, inguinal canal or abdominal area  160 unrelated control subjects All Pts & controls recruited from A.I. duPont Hospital for Children [Wilmington]  Exclusion criteria: boys with a hx of prematurity, and those with retractile testes, or that spontaneously descended within the first 6 months of life.												
<b>Sample and genetic analysis methods</b>														
30 SNPs were measured along the ESR1 gene. Leukocyte genomic DNA was amplified by PCR with primers designed to amplify a 300-bp region around e/SNP. Genotyping performed by 5' nuclease assay (TaqMan method). Haplotype analysis performed to identify one associated with cryptorchidism phenotype.	Germline DNA extracted from peripheral blood. For controls, DNA from frozen semen was used too. Five polymorphisms for the AGATA haplotype were genotyped by PCR protocols. Genotyping quality controls were performed.	Genomic DNA was obtained from peripheral blood leukocytes, buccal swab, or tissues from orchidopexy or circumcision. Puregene DNA purification kits were used for DNA extraction. TaqMan SNP Genotyping Assays was used to genotype the 5 SNPs (SNP10-SNP14). Haplotype association analysis performed using PLINK v1.01.												
<b>Cryptorchidism ascertainment</b>														
Not specified	Not specified	Not specified												
<b>Control selection</b>														
Controls (came from the urban or suburban area of Tokyo or Yamagata City) were found to have no discernible abnormality by cytogenetic, skeletal, and endocrine studies	Of the 168 men, 109 fathered at least one child spontaneously or had normal fertilization after IVF for pure female (tubal) factor infertility, whereas 59 men were students from the local university.	The 160 controls were healthy male infants and children presenting for circumcisions or voiding problems. Each boy had full descent of the testes, no significant testicular retractility, and no additional urogenital anomalies or known syndromes.												
<b>Haplotype frequency distribution and association</b>														
15 of 30 SNPs presented certain degree of heterozygosity. These were in HWE. Within block covering SNPs 10-14, four major haplotypes accounted for 95% (patients) and 93% (controls) of all haplotypes. AGATA haplotype frequency was greater in patients than in controls. SNP 12: A vs G OR=1.99 (95% CI 1.07, 3.67)	SNP10-13 did not deviate from HWE but SNP14 deviated in both cases and controls. SNP12 frequencies: <table border="1"> <thead> <tr> <th></th> <th>GG</th> <th>GA</th> <th>AA</th> </tr> </thead> <tbody> <tr> <td>Crypto</td> <td>101 (85.6)</td> <td>17 (14.4)</td> <td>0 (0.0)</td> </tr> <tr> <td>Control</td> <td>124 (73.8)</td> <td>43 (25.6)</td> <td>1 (0.6)</td> </tr> </tbody> </table> OR=0.50 (95% CI 0.28, 0.90)		GG	GA	AA	Crypto	101 (85.6)	17 (14.4)	0 (0.0)	Control	124 (73.8)	43 (25.6)	1 (0.6)	No difference in distribution of race was found between cases and controls. SNPs in Caucasian group were in HWE. No difference for SNP12 in Caucasians. Severe cases were more likely to have a GG genotype than moderate cases, OR=8.8 (1.1, 69.7). G allele OR=9.3 (1.2, 73.4). Mild cases were also less likely to have GG genotype and G allele compared to severe case but not significantly.
	GG	GA	AA											
Crypto	101 (85.6)	17 (14.4)	0 (0.0)											
Control	124 (73.8)	43 (25.6)	1 (0.6)											

## Appendix D - Evidence table of studies assessing the association between maternal alcohol consumption and cryptorchidism

Damgaard2007	Jensen2007	Strandberg-Larsen2009	Mongraw-Chaffin2008
<b>Study characteristics</b>			
<p>Copenhagen, Denmark 1997-00, Turku, Finland 1997-99. Prospective birth cohort study.</p> <p>Eligible women (2,229 Danish and 2,728 Finnish) residing in hospital referral areas were consecutively recruited during early pregnancy.</p> <p>Inclusion criteria: both parents and grandparents of the unborn child had to be born and raised in Denmark or Finland.</p>	<p>Aarhus, Denmark 1984-87. Prospective collection of prenatal exposures and obstetric information from a 16–19 years of follow-up in a nationwide patient register.</p> <p>Population-based random sample (n = 15,434) of boys born January 1984 and December 1987 drawn in the Danish Civil Registration system to assess comparability of the cohort.</p>	<p>Copenhagen, Denmark 1996–2002 (The Danish National Birth Cohort). Partial overlap (22%) with Damgaard’s Danish cohort.</p> <p>Pregnant women invited to participate by their GP at first antenatal visit. Fifty percent of GP participated and 60% of invited women consented. Analysis limited to women who gave birth to a singleton boy and completed first interview.</p>	<p>Berkeley, CA, USA 1959-1967 Prospective study ≥40-yr f-up of 20,754 pregnancies occurring b/w 1959-67 in CA.</p> <p>Cases were matched to three controls from the same cohort on birth year and race. If more than three controls were available, these were selected at random from the pool of matches.</p>
<b>Sample characteristics</b>			
<p>Denmark, 1,042 boys (1,029 mothers), Finland 1,454 boys (1,446 mothers).</p> <p>128 cryptorchid boys at birth [matched w/ 2,368 healthy controls] (94 Danish, 34 Finnish) Only 33 (19 Danish, 14 Finnish) [matched w/ 2,215 healthy controls] remained cryptorchid at 3 mos.</p>	<p>5,716 boys, 270 cases of crypto diagnosed, 185 of these underwent orchiopexy.</p> <p>Mothers: average gestational alcohol intake of five or more drinks per week. Age, parity, alcohol consumption smoking habits, infertility treatment.</p>	<p>41,268 live-born singleton, 1,598 cases. Of these, 355 boys were diagnosed with crypto but had no maternal report. Of the 810 boys diagnosed with crypto, [“diagnosis of crypto”] 398 had orchiopexy [“orchipexy”]</p>	<p>84 cases at birth persisting to at least age 2 yrs among 7,574 live-born sons whose mothers were interviewed in early pregnancy.</p>
<b>Alcohol consumption ascertainment</b>			
<p>Quantitative data on alcohol consumption, smoking, and caffeine intake obtained by questionnaire and/or interview once during the third trimester of pregnancy, before pregnancy outcome known.</p> <p>For a sub-group (n = 465) [no specification on who they were – cohort and/or controls], information on alcohol consumption was obtained twice during pregnancy by interviews.</p>	<p>Pregnant women attending their last scheduled regular antenatal care visit (~36<sup>th</sup> gestational wk), filled in a comprehensive self completed questionnaire, and returned it by mail.</p> <p>Alcohol/wk [units = 12g. of alcohol]. Binge episodes [intake of 8+ units of alcohol on one occasion]</p>	<p>Maternal alcohol consumption assessed in two computer-assisted telephone interviews around gestational weeks 17 and 32.</p> <p>One drink = one bottle of beer (0.33 l), one glass of wine, or one glass of spirits, each approx. 12 g of alcohol. Total = sum of weekly intake. Binge drinking: 5+ drinks on a single occasion since the onset of pregnancy.</p>	<p>In structured interview, mother’s alcohol consumption ascertained by asking how many glasses of beer, wine, or whiskey she drank in a week.</p> <p>Total drinks of alcohol consumed per wk were obtained by adding the number of glasses drunk per week for the 3 beverages.</p>
<b>Cryptorchidism ascertainment</b>			
<p>Examination at birth and 3 months. Preterm boys examined at expected delivery date. Gestational age based on routine US at 18–20 wks. When not available (2.1%), last menstrual period used. Data on birth weight and parity obtained from birth records.</p> <p>Testis considered cryptorchid if found in a high scrotal, supra-scrotal or inguinal position or if non-palpable. Retractable testes considered a normal variant.</p> <p>Majority were transient cryptorchid cases but still presented slightly elevated gonadotropin levels at 3 months.</p>	<p>Cases ascertained by clinical examination at birth and at 3 mos. Most cases were transient (spontaneous descent within 3 months). Among persistent cases no excess risk was observed (only 33 observations).</p> <p>All diagnosis or surgical procedures during hospital admissions of these boys in the yrs. 1984–2003 were extracted from the Danish National Patient Register.</p> <p>Two endpoints: boys having a crypto diagnosis without orchiopexy (‘no orchiopexy’) and boys with a diagnosis who also underwent</p>	<p>Maternal report at 6 or 18 months and/or crypto dx in National Hospital Discharge Registry (NHDR), 398 of 1,598 were verified by orchiopexy (identified in the NHDR).</p>	<p>Condition should be present at 2 yrs of age to examine risk factors for persistent crypto.</p>

Damgaard2007	Jensen2007	Strandberg-Larsen2009	Mongraw-Chaffin2008
orchiopepy ('orchiopepy').			
<b>Univariate estimate</b>			
Drinks/week continuous (*): (n= 2,477) 1.26 (1.13–1.40)	Risk ratios, ref (<1 drink/wk) a) All cases (n=270) 1-4 0.9 (0.7, 1.1) 5-9 0.7 (0.4, 1.3) ≥10 0.8 (0.2, 2.5)	Hazard ratio, ref (0 drinks/wk) a) Maternal report of crypto dx 0.5-1.5 0.87 2-3.5 0.82 4+ 0.63	Odds ratio 3.3 drinks/wk 1.05 (0.86, 1.29)
Binge [five or more alcoholic drinks at one occasion/ day] episodes: Yes 1.25 (0.85, 1.87)	b) Surgery (n=185) 1-4 0.8 (0.6, 1.1) 5-9 0.6 (0.3, 1.2) ≥10 0.7 (0.2, 3.0)	b) Crypto dx c) Surgery 0.5-1.5 0.92 0.89 2-3.5 0.87 0.77 4+ 0.61 0.77	
(*). Additionally, several OR from 0 drink to ≥ 9 drinks per week	c) No surgery (n=85) 1-4 1.0 (0.6, 1.5) 5-9 1.1 (0.5, 2.4) ≥10 0.9 (0.1, 6.7)	Binge drinking episodes a) Maternal report of crypto dx 1 0.90 2 0.89 3+ 0.99	
Drinks/wk OR (95% CI) 0 Ref 0.1-4.9 1.69 (1.16, 2.45) ≥ 5 4.97 (2.00, 12.40)	Binge episodes (Ref=No) RR for binge drinking a) All cases 1.2 (0.9, 1.7) b) Surgery 1.3 (0.9, 1.9) c) No surgery 1.0 (0.6, 1.8)	b) Crypto dx c) Surgery 1 0.86 0.74 2 0.79 0.56 3+ 0.94 0.93	
<b>Multivariate estimate</b>			
MV odds ratio (OR) adjusted for country, smoking, caffeine intake, maternal age, social class, parity, birth weight. Ref: 0 drinks/wk.	MV risk ratios (RR) adjusted for maternal and paternal age at delivery, time to index pregnancy, infertility txt, parity, SES group, mother's daily smoking, birth weight, gestational age, placental weight. Ref: < 1 drink/wk	MV hazard ratio (HR) adjusted for maternal age, parity, time to pregnancy, infertility txt, self- reported diabetes, smoking, occupational status. Ref: 0 drinks/wk	MV odds ratio (OR) adjusted for smoking, caffeine consumption, body mass index, son's birth weight.
Drinks/week continuous: (n=2,477) 1.17 (1.03, 1.34)	a) All cases (n=270) 1-4 0.9 (0.7, 1.2) 5-9 0.7 (0.4, 1.3) ≥10 0.7 (0.2, 2.2)	a) Maternal report of crypto dx 0.5-1.5 0.89 (0.79, 0.99) 2-3.5 0.82 (0.69, 0.99) 4+ 0.63 (0.41, 0.94)	3.3 drinks/wk 0.99 (0.83, 1.20)
Binge episodes: Yes 1.18 (0.77, 1.83)	b) Surgery (n=185) 1-4 0.8 (0.6, 1.2) 5-9 0.5 (0.2, 1.2) ≥10 0.6 (0.1, 2.7)	b) Crypto dx 0.5-1.5 0.93 (0.80, 1.08) 2-3.5 0.87 (0.68, 1.12) 4+ 0.58 (0.32, 1.06)	
Drinks/wk OR (95% CI) 0.1-4.9 0.95 (0.61, 1.49) ≥ 5 3.10 (1.05, 9.10)	c) No surgery (n=85) 1-4 1.0 (0.6, 1.6) 5-9 1.1 (0.5, 2.6) ≥10 0.8 (0.1, 6.0)	c) Surgery 0.5-1.5 0.91 (0.73, 1.13) 2-3.5 0.79 (0.54, 1.15) 4+ 0.77 (0.36, 1.63)	
	Binge episodes No (Ref) a) All cases (n=270) Yes 1.3 (0.9, 1.8)	Binge drinking a) Maternal report of crypto dx 1 0.88 (0.77, 1.02) 2 0.85 (0.70, 1.06) 3+ 0.94 (0.73, 1.21)	
	b) Surgery (n=185) Yes 1.4 (0.9, 2.1)	b) Crypto dx 1 0.83 (0.68, 1.01) 2 0.73 (0.53, 1.01) 3+ 0.85 (0.60, 1.21)	
	c) No surgery (n=85) Yes 1.0 (0.5, 1.8)	c) Surgery 1 0.70 (0.52, 0.94) 2 0.50 (0.29, 0.85) 3+ 0.81 (0.49, 1.34)	

## **Appendix E - Definitions**

*Pediatricians* are physicians who are board certified in pediatrics.

*Primary Care Providers*, in the context of these guidelines, are any licensed healthcare professionals who provide primary care services to patients.

*Providers* include, in the context of these guidelines, any licensed health care professional involved in the care of patients with suspected cryptorchidism such as: allied health professionals (Nurse Practitioners, Registered Nurses, Physicians' Assistants), physicians: pediatricians (including subspecialists), urologists and general surgeons (including subspecialists), or other licensed health professionals with specialized training and/or medical expertise in this area.

*Specialists* refer to a physician or allied health professional or other licensed health professional with specialized training/expertise, who are trained in a particular specialty of medicine including urology, endocrinology, or pediatric subspecialty.

*Surgical Specialists* refer to any physician who is trained or who has expertise in a specific branch of surgery, including (in the context of this guideline) urologic surgery, pediatric general surgery, or pediatric urology.

## Appendix F - Distribution of nonpalpable testes location as identified by diagnostic laparoscopy

Group	Absent/Unknown Vanishing/Atrophic	Inguinal	Internal ring/Peeping	Intra- abdominally/ Above internal ring
#1: ≤16 year-old <sup>68-100</sup>	Meta-analytic mean*=33.2%, I <sup>2</sup> =92%  Minimum=3.4%	Skewed distributio n with 44% null	Skewed distribution with 44% null	Meta-analytic mean*=40.3%, I <sup>2</sup> =91%  Minimum=8.8 %
#2: Prepubertal + older males or unknown <sup>101-115</sup>	Skewed distribution with meta-analytic mean*=32.5%, I <sup>2</sup> =92%  Minimum=13.5%	Skewed distributio n with 33% null	Zero- inflated distribution with 47% null. Non- null values centered at 35%.	Meta-analytic mean*=36.9%, I <sup>2</sup> =78%  Minimum=21. 6%

\*Meta-analytic mean indicates that estimate of the mean was obtained using meta-analysis. I<sup>2</sup> indicates the degree of heterogeneity.

These cohorts were divided into two groups (see Table 3). Group #1 includes boys 16 years-old and younger (34 cohorts).<sup>69-100</sup> Group #2 includes cohorts with a large proportion of prepubertal boys (mean or median age is small), as well as some cases where the male is older than 16.<sup>101-115</sup> Two cohorts in Group #2 do not report the age distribution.<sup>110,113</sup> Therefore, the distribution of the location of testes was estimated separately for these two sets of cohorts.

In Group #1, a total of 3,166 testes were assessed for an average of 93 testes per study; for Group #2, 1,525 testes were assessed and yielded an average of 102 testes per study. For the purpose of determining the distribution, absent/unknown testes and vanishing or atrophic were all combined because it is suspected that these are not well differentiated in some studies. Eight studies<sup>101-103,106,112,114,115</sup> showed differentiation between absent and vanishing testes; the average rate of absent/unknown testes in these cohorts was 3%.

Distributions were similar between the two groups, excluding older cohorts. In these cases, the minimum percent of intra-abdominally located testes was higher in Group #2 than in Group #1 (21.6% vs. 8.8%). Similarly, the minimum percent of absent or vanishing testes was also higher in Group #2 than in Group #1 (13.5% vs 3.4%). These data suggests more intra-abdominally testes will be present in older boys.

This meta-analysis shows that in any cohort of diagnostic laparoscopy for nonpalpable testes, at least 8% of testes will be located intra-abdominally and a small percent of testes will be absent/vanishing/atrophic (3.4%).

### References for Appendix F

Full citations available in *Evaluation and Treatment of Cryptorchidism: AUA Guideline*.







## Appendix H: Quality Assessment of Individual Studies

Quality was assessed using three published tools, depending on the study design. Two reviewers independently assessed the quality of each study, and then results were adjudicated. For randomized controlled trials (RCTs), we used the Cochrane Risk of Bias tool<sup>1</sup> for cohort studies, the Newcastle-Ottawa Quality Assessment Scale<sup>2</sup>; and for prognostic studies of imaging, the Quality Assessment of Diagnostic Accuracy Studies-Revised (QUADAS-2) tool<sup>3</sup>. The results of these tools were then translated to the Agency for Healthcare Research and Quality standard of “good,” “fair,” and “poor” quality designations using conversion thresholds developed by the team, as no explicit guidance exists.

Thresholds for converting the Newcastle-Ottawa scales to AHRQ standards (good, fair, and poor):

*Good quality:* 3 or 4 stars in selection domain AND 1 or 2 stars in comparability domain AND 2 or 3 stars in outcome/exposure domain

*Fair quality:* 2 stars in selection domain AND 1 or 2 stars in comparability domain AND 2 or 3 stars in outcome/exposure domain

*Poor quality:* 0 or 1 star in selection domain OR 0 stars in comparability domain OR 0 or 1 stars in outcome/exposure domain

## References for Appendix H

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<sup>1</sup> Higgins JPT, Altman DG, Sterne JAC. Chapter 8: Assessing risk of bias in included studies. In: Higgins JPT, Green S, eds. Cochrane Handbook for Systematic Reviews of Interventions Version 5.1.0 (updated March 2011). The Cochrane Collaboration; 2011.

<sup>2</sup> Wells GA, Shea B, O'Connell D, et al. The Newcastle-Ottawa Scale (NOS) for assessing the quality of nonrandomised studies in meta-analyses. [www.ohri.ca/programs/clinical\\_epidemiology/oxford.asp](http://www.ohri.ca/programs/clinical_epidemiology/oxford.asp). Accessed January 24, 2012.

<sup>3</sup> Whiting PF, Rutjes AW, Westwood ME, et al. QUADAS-2: a revised tool for the quality assessment of diagnostic accuracy studies. *Ann Intern Med*. 2011 Oct 18;155(8):529-36. PMID: 22007046.